

STA257 (Probability and Statistics I) Lecture Notes, Fall 2025

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Note: I will update these notes regularly, posting them on the course web page each evening after lectures (though without annotations). However, they are just rough, point-form notes, with no guarantee of completeness or accuracy. They should in no way be regarded as a substitute for attending and learning from all the lectures, studying the course textbook, and doing the suggested homework exercises.

Introduction

- Course Information: See the course web page at: probability.ca/sta257

- Register for PollEverywhere: probability.ca/sta257/pollinfo.html

→ **USE YOUR REGULAR UofT EMAIL!**

- Who here is doing a specialist or major program involving: Statistics / Data Science? Mathematics? Actuarial Science? Computer Science? Economics/Commerce? Physics/Chemistry/Biology? Education? Psychology/Sociology? Engineering? Other?

- Who here has seen probabilities in elementary school? high school? STA130?

→ Don't worry, we will start from scratch. (Just need math and logic.)

- Life is full of randomness and uncertainty: lotteries, card games, computer games, gambling, weather, TTC, airplanes, friends, jobs, classes, science, finance, elections, diseases, safety/risk, demographics, internet routing, legal cases, ... whenever we're not sure of the outcome or what will happen next.

- Lots of interesting probability questions to solve! Such as ...

→ What's the probability you'll win the Lotto Max jackpot, i.e. that you will choose the correct 7 distinct numbers between 1 and 50?

→ If 200 students each flip a fair coin, then how many Heads is the most likely? How likely? What's the probability of more than 150 Heads?

→ If you repeatedly roll a fair 6-sided die [show], then how many rolls will there be on average before the first time you roll a 3?

→ At a party of 40 people, what is the probability that some pair of them have the same birthday?

→ If a disease affects one person in a thousand, and a test for the disease has 99% accuracy, and you test positive, then what is the probability you have the disease?

→ If you pick a number uniformly at random between 0 and 1, then what is the probability that you pick exactly the number $3/4$?

→ Three-Card Challenge. [demonstration] What are the probabilities of the initial (front) colour? Then, what are the probabilities of the back colour?

- History of Mathematical Probability Theory (in brief):

→ Mathematics is very precise and certain. For thousands of years, it simply ignored the uncertainty of probabilities.

→ Then, in 1654, the French writer Antoine Gombaud (the “Chevalier de Méré”) asked the mathematician Pierre de Fermat some gambling questions:

→ Which is more likely (or are they the same) (and are they more than 50%):

(a) Get at least one six when rolling a fair six-sided die 4 times; or

(b) Get at least one pair of sixes when rolling two fair six-sided dice 24 times?

→ He thought (a) was $4 \times (1/6) = 2/3$, and (b) was $24 \times (1/36) = 2/3$. Correct?

→ Also: (c) Suppose a gambler is playing a best-of-seven match, where whoever wins 4 (fair) games first is the winner, and so far they have won 3 times and lost 1, but then the match gets interrupted. What is the probability that they would have won the match, if it had been allowed to continue?

→ Fermat then corresponded with the mathematician Blaise Pascal to find solutions to these questions (later!), and mathematical probability theory was born!

POLL: If you have independent probability $1/2$ of winning each game, and you are up 3 games to 1, what do you think is the probability that you will win 4 games first?
(A) $1/2$. **(B)** $2/3$. **(C)** $3/4$. **(D)** $7/8$. **(E)** No idea. [Best guess only – later.]

- So, can probabilities be studied mathematically?

→ Can we use certain mathematics to study the uncertainty of probabilities?

→ Yes! That’s why we’re here! To be certain about our uncertainty!

→ But we have to define our terms carefully ...

Sample Space (§1.2) (i.e. Section 1.2 of the textbook)

- The first part of any probability model is the **sample space**, written S , which is the set of all possible outcomes.

→ e.g. flip a coin: $S = \{\text{Heads, Tails}\}$, or $S = \{H, T\}$.

→ e.g. flip a coin three times in a row:

$S = \{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\}$.

→ Or, if we only care about the number of Heads: $S = \{0, 1, 2, 3\}$.

→ e.g. tonight’s dinner: $S = \{\text{Beef, Chicken, Fish}\}$. (Assume one.)

→ e.g. the number of bees I will see on my walk home: $S = \{0, 1, 2, 3, \dots\}$.

→ e.g. the price of IBM stock next month: $S = [0, \infty)$.

→ e.g. the height (in cm) of the next student I meet: $S = (0, \infty)$.

→ e.g. your grade in this class: $S = \{0, 1, 2, 3, \dots, 100\}$.

→ e.g. roll one six-sided die: $S = \{1, 2, 3, 4, 5, 6\}$.

→ e.g. roll two six-sided dice: $S = \{1, 2, 3, 4, 5, 6\}^2$, i.e.

$S = \{11, 12, 13, 14, 15, 16, 21, 22, 23, 24, 25, 26, 31, 32, 33, 34, 35, 36, 41, 42, 43, 44, 45, 46, 51, 52, 53, 54, 55, 56, 61, 62, 63, 64, 65, 66\}$.

→ Or, if we only care about the sum, instead maybe take $S = \{2, 3, 4, 5, 6, 7, 8, 9, 10, 11, 12\}$.



- e.g. “Pick any integer between 1 and 10”: $S = \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10\}$.
- e.g. “Pick any number between 0 and 1”: $S = [0, 1]$. (important case!)
- Summary: The sample space S can be any non-empty set which contains all of the possible outcomes. Simple!
- But it gets more interesting when we also have ...

Probabilities and Events (§1.2)

- An **event** A is “any” subset $A \subseteq S$.
- For any event A , we can define the **probability** $P(A)$ that it will occur.
 - e.g. flip a “fair” coin: $P(H) = P(T) = 1/2$.
 - (Note: We often use e.g. “ $P(H)$ ” as shorthand for “ $P(\{H\})$ ”, etc.)
 - e.g. roll a fair six-sided die: $P(1) = P(2) = P(3) = P(4) = P(5) = P(6) = 1/6$.
 - e.g. tonight’s dinner: maybe $P(\text{Beef})=0.40$, $P(\text{Chicken})=0.15$, and $P(\text{Fish})=0.45$.
 - (Note: We could also write $P(\text{Fish}) = 45\%$, etc. Usually percentages are good for intuition, but pure probabilities (not percentages) are better for calculation.)
 - e.g. flip three fair coins: $P(HHH) = P(HHT) = \dots = P(TTT) = 1/8$.
 - e.g. roll two fair dice: $P(11) = P(12) = \dots = P(65) = P(66) = 1/36$.
 - e.g. Pick any integer between 1 and 10. [Try it!]
- Could be “uniform”, i.e. $P(1) = P(2) = \dots = P(10) = 1/10$. Or instead, maybe ... $P(3)=P(6)=P(7)=0.2$, and $P(5)=0.1$, and $P(1)=P(2)=P(4)=P(8)=P(9)=P(10)=0.05$.
- e.g. Pick any number between 0 and 1, “uniformly” (“Uniform[0,1]”):
 $P([0, 1/2]) = 1/2$, $P([1/2, 1]) = 1/2$, $P([0, 1/3]) = 1/3$, $P([1/3, 2/3]) = 1/3$,
 and in general $P([a, b]) = b - a$ whenever $0 \leq a \leq b \leq 1$. Diagram:

- Or maybe instead $P([a, b]) = b^2 - a^2$ whenever $0 \leq a \leq b \leq 1$. Valid?
- Or maybe instead $P([a, b]) = (b - a)^2$ whenever $0 \leq a \leq b \leq 1$. Valid?

Basic Properties of Probabilities (§1.2)

- Let’s begin with a specific example (and then we will generalise):
- e.g. tonight’s dinner, with $P(\text{Beef})=0.40$, $P(\text{Chicken})=0.15$, and $P(\text{Fish})=0.45$.
 - Probability of Beef or Chicken = $P(\{\text{Beef}, \text{Chicken}\}) = P(\{\text{Beef}\}) + P(\{\text{Chicken}\}) = 0.40 + 0.15 = 0.55$.
 - Probability of any dinner = Probability of Beef or Chicken or Fish = $P(\{\text{Beef}, \text{Chicken}, \text{Fish}\}) = P(\{\text{Beef}\}) + P(\{\text{Chicken}\}) + P(\{\text{Fish}\}) = 0.40 + 0.15 + 0.45 = 1$.
 - Probability dinner is not Beef nor Chicken nor Fish = $P(\emptyset) = 0$.
- In general, certain properties must hold for any probability model (“axioms”):

- If A is an event, then $0 \leq P(A) \leq 1$.
- If $A = S$ is the event corresponding to all outcomes, then $P(A) = P(S) = 1$.
- Or, if $A = \emptyset$ is the event corresponding to no outcomes, then $P(A) = P(\emptyset) = 0$.
- **Additivity:** If A and B are disjoint events (i.e. $A \cap B = \emptyset$), e.g. $A = \{\text{Beef}\}$ and $B = \{\text{Chicken}\}$, then $P(A \cup B) = P(A) + P(B)$.
- More generally, if A_1, A_2, A_3, \dots are any sequence (finite or infinite) of disjoint events (i.e. $A_i \cap A_j = \emptyset$ whenever $i \neq j$), then $P\left(\bigcup_i A_i\right) = \sum_i P(A_i)$.
 - So, in particular, since $P(S) = 1$, all of the probabilities have to add up to 1.
 - e.g. $P(\text{Heads}) + P(\text{Tails}) = 0.5 + 0.5 = 1$.
 - e.g. $P(\text{Beef}) + P(\text{Chicken}) + P(\text{Fish}) = 0.40 + 0.15 + 0.45 = 1$.

Suggested Homework: 1.2.1, 1.2.2, 1.2.3, 1.2.4, 1.2.8, 1.2.9, 1.2.10, 1.2.11, 1.2.12, 1.2.13, 1.2.14, 1.2.15.

END WEDNESDAY #1

Derived Properties of Probabilities (§1.3)

- Once we know the above properties, then we can use them to prove others too:
- **Fact:** If A^C is the **complement** of A , i.e. the set of all outcomes which are not in A , then $P(A^C) = 1 - P(A)$. (Important! Remember this! Use this!)
 - Proof: Note that A and A^C are disjoint, so $P(A \cup A^C) = P(A) + P(A^C)$. But $P(A \cup A^C) = P(S) = 1$, so $1 = P(A) + P(A^C)$, i.e. $P(A^C) = 1 - P(A)$. ■
 - e.g. $P(\text{Fish}) = P(\text{not Beef or Chicken}) = 1 - P(\text{Beef or Chicken}) = 1 - 0.55 = 0.45$.
- **Fact:** For any events A and B , $P(A) = P(A \cap B) + P(A \cap B^C)$. (*)

Diagram:

→ Proof: The events $A \cap B$ and $A \cap B^C$ are disjoint, and $(A \cap B) \cup (A \cap B^C) = A$, so by additivity, $P(A \cap B) + P(A \cap B^C) = P(A)$. ■

→ e.g. integer between 1 and 10: $P(\text{even}) = P(\text{even and } \leq 4) + P(\text{even and } \geq 5) = P(\{2, 4\}) + P(\{6, 8, 10\})$.

• Re-arranging (*) also gives that: $P(A \cap B^C) = P(A) - P(A \cap B)$. (**)

• **Fact:** If $A \supseteq B$, then $P(A) = P(B) + P(A \cap B^C)$. (***)

→ Proof: This follows from (*), since if $A \supseteq B$, then $A \cap B = B$. ■

→ e.g. integer between 1 and 10: $P(\leq 7) = P(\leq 4) + P(\leq 7 \text{ but } \geq 5)$.

• **Monotonicity:** If $A \supseteq B$, then $P(A) \geq P(B)$. (Remember this!)

→ Proof: We must have $P(A \cap B^C) \geq 0$, so from (**),
 $P(A) = P(B) + P(A \cap B^C) \geq P(B) + 0 = P(B)$. ■

→ e.g. $P(\{\text{Beef}, \text{Chicken}\}) = 0.55 \geq 0.40 = P(\{\text{Beef}\})$.

• **Law of Total Probability – Unconditioned Version:** Suppose A_1, A_2, \dots are a sequence (finite or infinite) of events which form a partition of S , i.e. they are disjoint ($A_i \cap A_j = \emptyset$ for all $i \neq j$) and their union equals the entire sample space ($\bigcup_i A_i = S$), and let B be any event. Diagram:

Then $P(B) = \sum_i P(A_i \cap B)$. That is: $P(B) = P(A_1 \cap B) + P(A_2 \cap B) + \dots$

→ Proof: Since the $\{A_i\}$ are disjoint, and $A_i \cap B \subseteq A_i$, therefore the $\{A_i \cap B\}$ are also disjoint. Furthermore, since $\bigcup_i A_i = S$, therefore $\bigcup_i (A_i \cap B) = S \cap B = B$. Hence, $P(B) = P\left(\bigcup_i (A_i \cap B)\right) = \sum_i P(A_i \cap B)$. ■

→ e.g. integer between 1 and 10: Suppose $A_1 = \{\leq 4\} = \{1, 2, 3, 4\}$, and $A_2 = \{\geq 5\} = \{5, 6, 7, 8, 9, 10\}$, and $B = \{\text{even}\} = \{2, 4, 6, 8, 10\}$. Then $P(\text{even}) = P(\text{even and } \leq 4) + P(\text{even and } \geq 5)$, i.e. $P(\{2, 4, 6, 8, 10\}) = P(\{2, 4\}) + P(\{6, 8, 10\})$.

• **Principle of Inclusion-Exclusion:** $P(A \cup B) = P(A) + P(B) - P(A \cap B)$.

→ (Of course, if they're disjoint ($A \cap B = \emptyset$), then $P(A \cup B) = P(A) + P(B)$.)

→ Intuition: $P(A) + P(B)$ counts each element of $A \cap B$ twice, so we have to subtract one of them off.

→ Proof: The events $A \cap B$, and $A \cap B^C$, and $A^C \cap B$, are all disjoint, and their union is $A \cup B$. Diagram:

Hence, $P(A \cup B) = P(A \cap B) + P(A \cap B^C) + P(A^C \cap B)$.

But from (**), $P(A \cap B^C) = P(A) - P(A \cap B)$ and $P(A^C \cap B) = P(B) - P(A \cap B)$.

Hence, $P(A \cup B) = P(A \cap B) + [P(A) - P(A \cap B)] + [P(B) - P(A \cap B)]$
 $= P(A) + P(B) - P(A \cap B)$. ■

→ e.g. integer between 1 and 10: $P(\text{even or } \leq 4) = P(\text{even}) + P(\leq 4) - P(\text{even and } \leq 4) = P(\{2, 4, 6, 8, 10\}) + P(\{1, 2, 3, 4\}) - P(\{2, 4\})$.

→ Or, $P(\text{even or perfect square}) = P(\text{even}) + P(\text{perfect square}) - P(\text{even and perfect square}) = P(\{2, 4, 6, 8, 10\}) + P(\{1, 4, 9\}) - P(\{4\})$.

• Optional: A more general Inclusion-Exclusion formula is in **Challenge 1.3.10**.

• Now, $P(A \cap B) \geq 0$, so $P(A \cup B) = P(A) + P(B) - P(A \cap B) \leq P(A) + P(B)$. (!)

• **Subadditivity:** For any sequence of events A_1, A_2, \dots , not necessarily disjoint, we still always have $P(A_1 \cup A_2 \cup \dots) \leq P(A_1) + P(A_2) + \dots$

→ (Of course, it would be equal if they are disjoint.)

→ Proof (§1.7): Let $B_1 = A_1$, and $B_2 = A_2 \cap (A_1)^C$, and $B_3 = A_3 \cap (A_1 \cup A_2)^C$, and $B_4 = A_4 \cap (A_1 \cup A_2 \cup A_3)^C$, and so on. (That is, each new B_n is the part of A_n which is not already part of A_1, \dots, A_{n-1} .) Diagram:

Then the $\{B_i\}$ are disjoint by construction, and $\bigcup_i B_i = \bigcup_i A_i$.

[Formally, the above construction ensures that $\bigcup_{i=1}^n B_i = \bigcup_{i=1}^n A_i$ for each finite n . Then, in the infinite case, $\bigcup_{i=1}^\infty B_i = \bigcup_{n=1}^\infty (\bigcup_{i=1}^n B_i) = \bigcup_{n=1}^\infty (\bigcup_{i=1}^n A_i) = \bigcup_{i=1}^\infty A_i$.] Also $B_i \subseteq A_i$ so $P(B_i) \leq P(A_i)$. Hence, $P(A_1 \cup A_2 \cup \dots) = P(B_1 \cup B_2 \cup \dots) = P(B_1) + P(B_2) + \dots \leq P(A_1) + P(A_2) + \dots$ ■

→ Alternative proof (for a finite number of events): Use induction! For $n = 2$ events, this follows from Inclusion-Exclusion. Then for $n \geq 3$ events, $P(A_1 \cup \dots \cup A_n) = P((A_1 \cup \dots \cup A_{n-1}) \cup A_n)$, which by Inclusion-Exclusion is $\leq P(A_1 \cup \dots \cup A_{n-1}) + P(A_n)$, which by induction is $\leq (P(A_1) + \dots + P(A_{n-1})) + P(A_n)$. ■

→ e.g. integer between 1 and 10: $P(\text{even or } \leq 4) \leq P(\text{even}) + P(\leq 4)$, i.e. $P(\{1, 2, 3, 4, 6, 8, 10\}) \leq P(\{2, 4, 6, 8, 10\}) + P(\{1, 2, 3, 4\})$.

[Note that we do not have “uncountable” subadditivity, e.g. for uniform on $S = [0, 1]$, if $A_x = \{x\}$ for each $x \in S$, then $P(\bigcup_{x \in S} A_x) = P(S) = P([0, 1]) = 1$, even though $P(A_x) = P(\{x\}) = 0$ for each individual $x \in S$, so also $\sum_{x \in S} P(A_x) = \sum_{x \in S} 0 = 0$.]

Suggested Homework: 1.3.1, 1.3.2, 1.3.3, 1.3.4, 1.3.5, 1.3.7, 1.3.8, 1.3.9.

Uniform Probabilities on Finite Spaces (§1.4)

• Suppose $S = \{s_1, s_2, \dots, s_n\}$ is some finite sample space, of finite size $|S| = n$, and each element is equally likely.

→ Then $P(s_1) = P(s_2) = \dots = P(s_n) = 1/n$. (“discrete uniform distribution”)

→ And for any event $A = \{a_1, a_2, \dots, a_k\}$, by additivity we have

$$P(A) = P(a_1) + P(a_2) + \dots + P(a_k) = \frac{1}{n} + \frac{1}{n} + \dots + \frac{1}{n} = \frac{k}{n} = \frac{|A|}{|S|}.$$

→ So, in this case, we just need to count the number of elements in A , and divide that by the number of elements in S . Easy!?! Sometimes!

• e.g. Roll a fair six-sided die. What is $P(\geq 5)$?

→ Here $S = \{1, 2, 3, 4, 5, 6\}$ so $|S| = 6$. All equally likely.

→ Also $A = \{5, 6\}$ so $|A| = 2$.

→ So, $P(\geq 5) = P(A) = |A| / |S| = 2/6 = 1/3$. Easy!

• Flip two fair coins. What is $P(\# \text{ Heads} = 1)$?

POLL: (A) 1/4. (B) 1/3. (C) 1/2. (D) 3/4. (E) 1. (F) No idea.

→ Here $S = \{HH, HT, TH, TT\}$, all equally likely. So, $|S| = 4$.

→ And, $A = \{HT, TH\}$. So, $|A| = 2$.

→ Hence, $P(A) = |A| / |S| = 2/4 = 1/2$. Easy!

- e.g. Roll one fair six-sided die, and flip two fair coins.

What is $P(\# \text{ Heads} = \text{Number Showing On The Die})$? (Best guess?)

POLL: (A) 1/6. (B) 1/8. (C) 1/12. (D) 1/16. (E) 1/24. (F) No idea.

→ Here $S = \{1HH, 1HT, 1TH, 1TT, 2HH, \dots, 6TT\}$. All equally likely.

→ But what is $|S|$?

→ **Multiplication Principle:** If S is made up by choosing one element of each of the subsets S_1, S_2, \dots, S_k , i.e. if $S = S_1 \times S_2 \times \dots \times S_k$, then what is $|S|$? Well, \dots
 $|S| = |S_1| |S_2| \dots |S_k|$.

→ In our example, $S_1 = \{1, 2, 3, 4, 5, 6\}$, and $S_2 = \{H, T\}$, and $S_3 = \{H, T\}$, so
 $|S| = |S_1| |S_2| |S_3| = 6 \cdot 2 \cdot 2 = 24$.

→ And what about A ? Well, think about the possibilities \dots

$A = \{1HT, 1TH, 2HH\}$. (No other combination works. Why?) So, $|A| = 3$.

→ Hence, $P(\# \text{ Heads} = \text{Number Showing On The Die}) = |A| / |S| = 3/24 = 1/8$.

→ [Alternatively (later): $(1/6)(1/2) + (1/6)(1/4) = (1/12) + (1/24) = 3/24 = 1/8$.]

- e.g. Roll three fair six-sided dice. What is $P(\text{sum} \geq 17)$?

→ Here $S = \{1, 2, 3, 4, 5, 6\}^3$ so $|S| = 6^3 = 216$. All equally likely.

→ But what is A ? Think about it \dots

Here $A = \{666, 566, 656, 665\}$ (why?), so $|A| = 4$.

→ So, $P(\text{sum} \geq 17) = P(A) = |A| / |S| = 4/216 = 1/54$.

→ **Exercise:** What about $P(\text{sum} \geq 16)$? $P(\text{sum} \geq 15)$?

- **Chevalier de Méré's historical 1654 questions:**

- (a) What is $P(\text{get at least one six when rolling a fair six-sided die 4 times})$?

→ Here $S = \{1, 2, 3, 4, 5, 6\}^4$, so $|S| = 6^4 = 1296$. All equally likely.

→ And what is $|A|$? Tricky. Easier to consider \dots

→ $A^C = \{\text{no sixes in four rolls}\} = \{1, 2, 3, 4, 5\}^4$, so $|A^C| = 5^4 = 625$.

→ So, $P(A^C) = |A^C| / |S| = 5^4 / 6^4 = 625 / 1296 \doteq 0.482$.

→ So, $P(A) = 1 - P(A^C) \doteq 1 - 0.482 = 0.518$. More than 50%.

→ (Alternatively: By "independence" [later], $P(A) = 1 - (5/6)^4 \doteq 0.518$.)

- (b) What is $P(\text{get at least one pair of sixes when rolling a pair of fair six-sided dice 24 times})$?

→ Here $S = \left(\{1, 2, 3, 4, 5, 6\}^2\right)^{24}$, so $|S| = (6^2)^{24} = 6^{48} (>10^{37})$. All equally likely.

→ And what is $|A|$? Tricky. Again, easier to consider \dots

→ $A^C = \{\text{no pair of sixes in 24 rolls}\} = \{11, 12, 13, \dots, 64, 65\}^{24}$, so $|A^C| = 35^{24}$.

→ So, $P(A^C) = |A^C| / |S| = 35^{24} / 6^{48} \doteq 0.509$.

→ So, $P(A) = 1 - P(A^C) \doteq 1 - 0.509 = 0.491$. Less than 50%.

→ (Again, alternatively by independence [later], $P(A) = 1 - (35/36)^{24} \doteq 0.491$.)

Suggested Homework: 1.4.1, 1.4.9, 1.4.10, 1.4.11, 1.4.12, 1.4.13.

END MONDAY #1
